# An anthropometry-based equation of fat mass percentage as a valid discriminator of obesity 

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#### Abstract

Objective: To develop a new predictive equation for fat mass percentage (\%FM) based on anthropometric measurements and to assess its ability to discriminate between obese and non-obese individuals. Design: Cross-sectional study. Setting: Mexican adults. Participants: Adults ( $n$ 275; 181 women) aged 20-63 years with BMI between 17.4 and $42 \cdot 4 \mathrm{~kg} / \mathrm{m}^{2}$. Results: Thirty-seven per cent of our sample was obese using \%FM measured by air-displacement plethysmography (BOD POD ${ }^{\circledR}$; Life Measurement Instruments). The fat mass was computed from the difference between weight and fat-free mass (FFM). FFM was estimated using an equation obtained previously in the study from weight, height and sex of the individuals. The \%FM estimated from the obtained FFM showed a sensitivity of $90 \cdot 3(95 \%$ CI $86 \cdot 8,93 \cdot 8) \%$ and a specificity of $58.0(95 \%$ CI $52 \cdot 1,63 \cdot 8) \%$ in the diagnosis of obesity. Ninety-three per cent of participants with obesity and $65 \%$ of participants without obesity were correctly classified. Conclusions: The anthropometry-based equation obtained in the present study could be used as a screening tool in clinical and epidemiological studies not only to estimate the $\% \mathrm{FM}$, but also to discriminate the obese condition in populations with similar characteristics to the participant sample.


Keywords Obesity BMI Body composition Fat mass Prediction equations Discriminant analysis

BMI is an indicator used internationally to classify obesity. Adults with BMI equal to or greater than $30.0 \mathrm{~kg} / \mathrm{m}^{2}$ are considered obese, regardless of age, sex or ethnicity. Despite being used widely, mainly due to ease of calculation, this approach is inconsistent with the definition of obesity as an excess of body fat that can be harmful to health ${ }^{(1)}$

Several authors have reported a positive association between BMI and fat mass percentage (\%FM) ${ }^{(2,3)}$; however, variation in fat mass between individuals is wide and BMI does not consider it. Individuals with the same BMI value can have differences in their \%FM. For example, in a sample of adult men, Smalley et al. found for a BMI value of $27.0 \mathrm{~kg} / \mathrm{m}^{2}$ that $\% \mathrm{FM}$ can range from 10 to $31 \cdot 7^{(4)}$.

The worldwide increased prevalence of obesity in recent years ${ }^{(5)}$, Mexico included ${ }^{(6)}$, has resulted in high rates of diseases associated with obesity ${ }^{(7-9)}$ and high costs to health services ${ }^{(10)}$. Because of this is important to have more precise adiposity measurements.

The use of a single BMI standard for both men and women cannot be justified on the basis of weight-height relationships. In most populations, BMI is dependent on height; weight does not universally vary with the square of height; and the relationship between weight and height differs significantly between males and females ${ }^{(11)}$. Furthermore, the same BMI cut-off of $30.0 \mathrm{~kg} / \mathrm{m}^{2}$ corresponds to different $\% \mathrm{FM}$ in the diagnosis of obesity for Caucasian men and women ( 25 or $35 \%$, respectively) ${ }^{(12)}$.

Excess adiposity is the main phenotypic feature that defines human obesity and has a pathophysiological role in most chronic diseases. Although BMI and waist circumference have been widely used to define obesity and central obesity, they do not represent body fat mass and fat distribution precisely ${ }^{(13)}$. BMI and waist circumference perform similarly as indicators of body fatness and are more closely related to each other than with percentage of body fat. These variables may be an
inaccurate measure of body fat percentage for an individual ${ }^{(14)}$. Some authors have highlighted the importance of studying fat mass. For example, Lee et al. ${ }^{(15)}$ showed a strong positive monotonic association between predicted fat mass and all-cause mortality; they found that, compared with those in the lowest fifth of predicted fat mass, men in the highest fifth had a hazard ratio of 1.35 ( $95 \%$ CI $1.26,1.46$ ) for mortality from all causes. Measuring the amount of fat mass is a central issue of studying obesity at the individual and population levels ${ }^{(16)}$.

Although there are various methods and instruments to measure the fat component of an individual ${ }^{(17)}$, most of them are expensive and non-viable for population studies.

Prediction equations of body composition have emerged as a cheaper, simple and more viable option. However, some of them were conceived to be used for bioimpedance ${ }^{(18,19)}$. The equations for adults that consider anthropometric measurements were mainly developed for specific populations and age groups ${ }^{(20-27)}$ (Table 1). This issue has not been addressed by a simple anthropometrybased model for Mexican adults.

The present study had two aims: (i) to develop and validate a fat-free mass (FFM) anthropometry-based equation to estimate $\% \mathrm{FM}$ in a sample of adults; and (ii) to assess its discriminant ability for obesity compared with the \%FM obtained by air-displacement plethysmography.

## Methods

## Participants

Two hundred and seventy-five volunteers aged 20-63 years from the Hidalgo State attended the plicometry laboratory of the Institute of Health Sciences of the Autonomous University of Hidalgo State located in Pachuca, Mexico. None of the individuals had a disease or physical condition that might affect their body volume such as oedema, dehydration, pregnancy or amputations, nor a phobia to confined places.

The inclusion criteria included a 12 h fast, no consumption of alcohol for 24 h before measurements, no strenuous exercise within 12 h before the measurements, avoidance of use of moisturizing lotions, no taking a shower within 6 h before the measurements, and women should not be in their menstrual period or lactating. Each participant completed a test with the air-displacement plethysmography unit (BOD POD ${ }^{\circledR}$ ) and a set of anthropometric measurements was taken on the same day and within the same hour.

The Faculty of Medicine Ethics Committee from the National Autonomous University of Mexico approved the protocol. All participants provided written informed consent before participation.

## Anthropometry

Standing height was measured with no shoes using a stadiometer (SECA, Hamburg Germany) to the nearest $0 \cdot 1$ cm . Body weight was measured in light clothing with no shoes using the BOD POD's digital weight scale. Waist circumference was measured to the nearest 0.1 cm midway between the lower costal margin and the iliac crest while the participant was in a standing position and at minimal respiration; it was performed using a flexible and inelastic measuring tape (Rosscraft Innovations Incorporated, Canada). The thickness of two skinfolds (triceps and subscapular) was measured using a Harpenden skinfold calliper (British Indicators, Burgess Hill, UK) and calculated according to Durnin and Womersley ${ }^{(28)}$. The sagittal abdominal diameter was measured to the nearest 0.1 cm after a normal exhalation while the participant was in a supine position on a firm examination table; the measurement was taken at the umbilicus level using the Hol-tain-Kahn abdominal calliper (Holtain Ltd, Crymych, UK), which is a portable sliding-beam calliper. All measurements were taken on the right side of the body, with each one repeated twice, by well-trained personnel. BMI was calculated as [weight $(\mathrm{kg})] /[\text { height }(\mathrm{m})]^{2}$ and was classified according to WHO standards ${ }^{(1)}$.

## Fat mass

We used the air-displacement plethysmography (ADP) method to assess $\% \mathrm{FM}_{\mathrm{ADP}}$ and $\mathrm{FFM}_{\mathrm{ADP}}$. The ADP equipment was the BOD POD ${ }^{\circledR}$ (Body Composition System manufactured by Life Measurement Instruments, Concord, CA, USA). Fields et al. ${ }^{(29)}$ validated the BOD POD as a reliable procedure to evaluate body composition in a wide range of population types, even those difficult to measure such as the elderly, children and individuals with obesity. The BOD POD is considered a reference method and has previously been described in detail ${ }^{(30)}$. Before measurement, the equipment was calibrated as recommended by the manufacturer. The participant entered the BOD POD wearing a tight-fitting swimsuit, a swim cap and without any jewellery. After the calibration procedure, the participant's body volume was measured while the participant was seated quietly in an erect posture in the test chamber and breathing normally; the participant was also instructed to stand with his/her hands on his/her thighs and his/her feet placed on the floor of the device. Thoracic gas volume was predicted by the BOD POD and the \%FM was derived by using Siri's formula for all participants. This equation was based on the Minnesota 'reference man', characterized by density ( $d_{0}=1.063 \mathrm{~g} / \mathrm{cm}^{3}$ ), fat ( $f_{0}=0.14$ ), water $\left(w_{0}=0.61\right)$ protein $\left(p_{0}=0.19\right)$ and mineral $\left(m_{0}=0.06\right)^{(31)}$.

## Statistical analysis

The sample was divided into two random groups, which resulted in a development sample and a validation sample. Descriptive statistics were used to characterize the

Table 1 Equations for estimation of fat mass/fat-free mass using anthropometrics in adults

| Study, year | Ethnicity | Age (years) | $n$ (sex) | Proposed model | SEE | $R^{2}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $\begin{gathered} \text { Hassager } \\ \text { et al }{ }^{(21)}, \\ 1986 \end{gathered}$ | Danish | 20-72 | $\begin{aligned} & 98(\mathrm{M}) \\ & 130(\mathrm{~F}) \end{aligned}$ | $\begin{aligned} & \text { FFM }(\mathrm{kg})=11+0.54 \mathrm{~W}-0.24 \mathrm{~S}+0.10 \mathrm{~A}(\mathrm{M}) \\ & \text { FFM }(\mathrm{kg})=17+0.67 \mathrm{~W}-0.20 \mathrm{~S}+0.03 \mathrm{~A}(\mathrm{~F}) \end{aligned}$ | $\begin{aligned} & 2.9 \mathrm{~kg} \\ & 2 \cdot 1 \mathrm{~kg} \end{aligned}$ | $\begin{aligned} & 0.76 \\ & 0.83 \end{aligned}$ |
| $\begin{aligned} & \text { Kwok } \\ & \text { et al. }{ }^{(22)} \text {, } \\ & 2000 \end{aligned}$ | Chinese | $\begin{aligned} & 69-79 \\ & 69-82 \end{aligned}$ | $\begin{aligned} & 261 \text { (M) } \\ & 352 \text { (F) } \end{aligned}$ | $\begin{aligned} & \% \mathrm{FM}=-27 \cdot 149+6 \cdot 137 \mathrm{Sex}+1 \cdot 12 \mathrm{BMI}+17 \cdot 308 \log (\mathrm{TSF}+\mathrm{BSF}) \\ & \text { where } \text { Sex }=1 \text { for male and } 2 \text { for female } \end{aligned}$ | 4.1 \%W | 0.81 |
| $\begin{aligned} & \text { Huerta } \\ & \text { et al. }{ }^{(23)} \text {, } \\ & 2007 \end{aligned}$ | Mexican | $>60$ | $\begin{aligned} & 100 \text { (M) } \\ & 102 \text { (F) } \end{aligned}$ | $\begin{aligned} & \mathrm{FM}(\mathrm{~kg})=-13171+0.165 \mathrm{CSF}+0.355 \mathrm{BSF}+0.521 \mathrm{~W}- \\ & 6.054 \mathrm{Sex} \\ & \text { where Sex }=1 \text { for male and } 0 \text { for female } \end{aligned}$ | 3.2 kg | 0.85 |
| $\begin{aligned} & \text { O'Connor } \\ & \text { et al al }{ }^{(24)} \text {, } \\ & 2010 \end{aligned}$ | Non-Hispanic White (37\%) Hispanic (29\%) AfricanAmerican (34\%) | 17-35 | $\begin{aligned} & 420 \text { (M) } \\ & 705 \text { (F) } \end{aligned}$ | $\begin{aligned} & \% \mathrm{FM}=-5.832+0.195 \Sigma 3 \mathrm{~S}-0.0005 \Sigma 3 \mathrm{~S}^{2}+0.608 \mathrm{BMI}- \\ & 2.399 \mathrm{AA}+0.053 \mathrm{H}(\mathrm{M}) \\ & \% \mathrm{FM}=1.260+0.169 \Sigma 3 \mathrm{~S}-0.0007 \Sigma 3 \mathrm{~S}^{2}+0.849 \mathrm{BMI}- \\ & 1.338 \mathrm{AA}+1.886 \mathrm{H}(\mathrm{~F}) \end{aligned}$ | $\begin{aligned} & 3.12 \% \\ & 3.64 \% \end{aligned}$ | NA NA |
| $\begin{aligned} & \text { Sandhu } \\ & \text { et al. }{ }^{(25)} \text {, } \\ & 2010 \end{aligned}$ | Indian | 20-59 | $\begin{aligned} & 58(\mathrm{M}) \\ & 60(\mathrm{~F}) \end{aligned}$ | $\begin{aligned} & \mathrm{FM}(\mathrm{~kg})=8 \cdot 46+0 \cdot 32 \mathrm{~W}-15 \cdot 16 \mathrm{~S}+9.54 \log (\Sigma 4 \mathrm{SF})(\mathrm{M}) \\ & \mathrm{FM}(\mathrm{~kg})=-20 \cdot 22+0.33 \mathrm{~W}+3.44 \mathrm{~S}+7 \cdot 66 \log (\Sigma 4 \mathrm{SF})(\mathrm{F}) \end{aligned}$ | $\begin{aligned} & 3.42 \mathrm{~kg} \\ & 3.01 \mathrm{~kg} \end{aligned}$ | $\begin{aligned} & 0.53 \\ & 0.72 \end{aligned}$ |
| $\begin{aligned} & \text { Hastuti } \\ & \text { et al }{ }^{(26)}, \\ & 2013 \end{aligned}$ | Javanese | 18-65 | 292 (M) | $\% F M=8.000+0.402 A S F+0.486 T S F+0.059 A$ | 3.68 \%BF | 0.69 |
| $\begin{aligned} & \text { Lee et al. }{ }^{(15)} \text {, } \\ & 2018 \dagger \end{aligned}$ | White Black | $\geq 18$ | $\begin{aligned} & 5239 \text { (M) } \\ & 4519 \text { (F) } \end{aligned}$ | $\begin{gathered} \text { FFM }(\mathrm{kg})=-14.729-0.071 \mathrm{~A}+0.210 \mathrm{~S}+0.468 \mathrm{~W}- \\ 0.441 \mathrm{Mex}+0.320 \mathrm{H}+1.821 \mathrm{~B}-0.784 \mathrm{OR}(\mathrm{M}) \end{gathered}$ | 2.96 kg | 0.88 |
|  | Mexican- <br> American Hispanic |  |  | $\begin{aligned} & \text { FFM }(\mathrm{kg})=-1.401-0.010 \mathrm{~A}+0.100 \mathrm{~S}+0.632 \mathrm{~W}- \\ & 0.225 \mathrm{WC}+0.315 \mathrm{AC}+0.091 \mathrm{CC}+0.040 \mathrm{TC}-0.304 \mathrm{TSF}- \\ & 0.021 \mathrm{SSF}+0.120 \mathrm{Mex}+0.097 \mathrm{H}+0.463 \mathrm{~B}-0.661 \mathrm{OR}(\mathrm{M}) \end{aligned}$ | 2.11 kg | 0.94 |
|  | Other |  |  | $\begin{aligned} & \text { FFM }(\mathrm{kg})=-14.292-0.046 \mathrm{~A}+0.201 \mathrm{~S}+0.347 \mathrm{~W}-0.448 \mathrm{Mex} \\ & -0.047 \mathrm{H}+1.128 \mathrm{~B}-0.384 \mathrm{OR}(\mathrm{~F}) \end{aligned}$ | 2.39 kg | 0.85 |
|  |  |  |  | $\begin{aligned} & \text { FFM }(\mathrm{kg})=-9.193-0.045 \mathrm{~A}+0.158 \mathrm{~S}+0.410 \mathrm{~W}- \\ & 0.040 \mathrm{WC}+0.095 \mathrm{AC}+0.193 \mathrm{CC}-0.105 \mathrm{TC}-0.152 \mathrm{TSF}- \\ & 0.004 \mathrm{SSF}-0.306 \mathrm{Mex}+0.082 \mathrm{H}+1.235 \mathrm{~B}-0.196 \mathrm{OR}(\mathrm{~F}) \end{aligned}$ | 2.22 kg | 0.87 |
| Aristizabal et al. ${ }^{(27)}$, | Colombian | 18-59 | 151 (F) | $\begin{aligned} & \% F M=11.76+0.324 T S F+0.133 C S F+0.347 A b C+0.068 \mathrm{~A}- \\ & 0.135 \mathrm{~S} \end{aligned}$ | $3 \cdot 12 \%$ | 0.72 |
| 2018 |  |  |  | \%FM $=27.39+0.264 W+0.381 \mathrm{AbC}-0.279 \mathrm{~S}$ | $3.44 \%$ | $0 \cdot 66$ |

SEE, standard sample error; $R^{2}$, regression coefficient; M, male; F, female; FFM, fat-free mass; W, weight (kg); S, stature (cm); A, age (years); \%FM, fat mass percentage; TSF, triceps skinfold (mm); BSF, biceps skinfold (mm); FM, fat mass; CSF, calf skinfold (mm); $\Sigma 3 S$, sum of three sex-specific skinfold sites (mm); AA, African-American; H, Hispanic; $\Sigma 4 S F$, sum of four skinfold sites; ASF, abdominal skinfold (mm); Mex, Mexican; B, Black; OR, Other race; WC, waist circumference ( cm ); AC, arm circumference ( cm ); CC, calf circumference ( cm ); TC, thigh circumference ( cm ); SSF, subscapular skinfold ( mm ); AbC, abdominal circumference (cm); \%W, percentage of weight; \%BF, percentage of body fat; NA, not available.
$\dagger$ Two equations were selected for each sex from the study of Lee et al. ${ }^{(15)}$.
anthropometric and body composition variables, and are reported as means and sD. Homogeneity of variances was tested by the sD test. Differences between the two samples were tested using the $\chi^{2}$ test for dichotomous variables and Student's $t$ test for continuous variables. Correlation between explanatory variables was calculated.

Normality of continuous variables was evaluated through the Shapiro-Wilk test and Q-Q plot. Stepwise and lasso multiple regressions were used to develop the model for estimating FFM. Eight variables were considered of interest and entered in the initial model: age, sex, body weight, height, sagittal abdominal diameter, waist circumference and the two skinfolds.

Several models were obtained from this process; the most parsimonious model with optimal regression coefficient ( $R^{2}$ ) and lowest standard sample error (SEE) was chosen. The simplicity of the predictor variables was
considered; that is, the ease of measurement considering the technique and instrument. Homoscedasticity and normal distribution of multiple regression residuals were verified. Although multicollinearity does not impact the predictive power of the model, it affects parsimony and therefore it was decided to develop a model with low multicollinearity. The variance inflation factor (VIF) was used to assess multicollinearity. Bland-Altman ${ }^{(32)}$ plots were created in the validation sample to determine levels of agreement between predicted $\left(\mathrm{FFM}_{\mathrm{EQ}}\right)$ and true FFM ( $\mathrm{FFM}_{\mathrm{ADP}}$ ). Although this technique of validation in a subsample is the most used in studies whose purpose is the development of prediction equations of some body component, we also did a leave-one-out cross-validation.
For the second study objective, once validation tests were completed, we used the discriminant multivariate technique to assess the ability of the \%FM estimated from the obtained FFM equation to classify obese and non-

Table 2 Anthropometry and other characteristics of the sample of Mexican adults ( $n 275$; 181 women) aged 20-63 years with BMI between 17.4 and $42.4 \mathrm{~kg} / \mathrm{m}^{2}$

|  | Development sample ( $n$ 140) |  |  | Validation sample ( $n$ 135) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Mean | SD | Range | Mean | SD | Range |
| Men† (\%) | 31.4 | - | - | 37.0 | - | - |
| Age (years) | $40 \cdot 0$ | 11.6 | 20-63 | 39.5 | $12 \cdot 2$ | 20-63 |
| Height (cm) | 159.6 | 9.6 | 141.4-184.5 | $160 \cdot 7$ | 8.7 | 141.9-184.8 |
| Weight (kg) | 73.9 | 12.6 | 44.5-109.1 | 74.4 | 13.2 | 45-108.5 |
| BMI (kg/m ${ }^{\text {2 }}$ ) | 29.0 | $4 \cdot 1$ | 17.4-42.4 | 28.7 | $4 \cdot 3$ | 17.8-39.5 |
| WC (cm) | 96.7 | 11.0 | 66.7-131.2 | 96.2 | 11.3 | 65.3-122.5 |
| TSF (mm) | 18.9 | 6.9 | 6.5-46.9 | $17 \cdot 5$ | 6.0 | 5.8-33.6 |
| SSF (mm) | 24.3 | 7.7 | 7.1-44.7 | 23.2 | $7 \cdot 6$ | 4.1-42.8 |
| SAD (cm) | 24.2 | 3.9 | 13.5-36.4 | 24.3 | 3.6 | 14.1-34.8 |
| FFM ${ }_{\text {ADP }}$ (kg) | $47 \cdot 6$ | 9.9 | 30.9-74.6 | 48.5 | 9.7 | 31.7-76.4 |
| $\mathrm{FM}_{\text {ADP }}(\mathrm{kg})$ | 26.3 | $8 \cdot 0$ | 8.6-55.1 | 25.9 | 8.9 | 3.55-46.6 |
| \%FM ${ }_{\text {ADP }}$ (\%) | 35.5 | $8 \cdot 1$ | 15.5-57.9 | 34.4 | 9.2 | 6.3-52.3 |

WC, waist circumference; TSF, triceps skinfold; SSF, subscapular skinfold; SAD, sagittal abdominal diameter; FFM ${ }_{\text {ADP }}$, fat-free mass as estimated from air-displacement plethysmography; FM graphy; \%FM ${ }_{\text {ADP }}$, fat mass percentage as estimated from air-displacement plethysmography.
$\dagger$ Differences between the samples were tested using the $x^{2}$ test for dichotomous variables and Student's $t$ test for continuous variables; no significant differences were found.
obese people. Sensitivity, specificity and predictive values of obesity were also calculated. Participants were considered obese if they had $\mathrm{BMI} \geq 30.0 \mathrm{~kg} / \mathrm{m}^{2}$. The cutoff points used were those recommended by the WHO to define obesity from $\%$ FM ( $\geq 35 \%$ for women and $\geq 25 \%$ for men) and to define overweight from $\% \mathrm{FM}$ (between 30 and $34 \%$ for women and between 20 and $24 \%$ for men). Finally, the \%FM was calculated as [(weight FFM)/weight] $\times 100$, where FFM was predicted from the obtained equation in the present study. Statistical analyses were performed using the statistical software package Stata ${ }^{\circledR}$ version 13.

## Results

The development sample included 140 participants (ninety-six women) aged between 20 and 63 years with BMI between 17.4 and $42.4 \mathrm{~kg} / \mathrm{m}^{2}$. The validation sample included 135 participants (eighty-five women) aged between 20 and 63 years with BMI between 17.8 and $39.5 \mathrm{~kg} / \mathrm{m}^{2}$. Both samples did not differ significantly on any of the anthropometric and body composition variables (Table 2).

Seventy-seven per cent of women and $70 \%$ of men were classified as obese using $\% \mathrm{FM}_{\mathrm{ADP}} ; 15$ and $18 \%$ of women and men, respectively, were classified as overweight. By contrast, when the classification was done from BMI, the prevalence of obesity dropped to about $40 \%$ in women and $32 \%$ in men (Table 3).

Multiple linear regression models obtained from the development sample showed that weight, height and sex were significantly associated with $\mathrm{FFM}_{\mathrm{ADP}}$ (Table 4). Models obtained using stepwise forward regression and lasso regression did not showed significant differences in the values of coefficients, $R^{2}$ and SEE.

The model chosen was the one formed from simple anthropometric variables:

$$
\begin{aligned}
\mathrm{FFM} & =-46 \cdot 03+0 \cdot 44 \mathrm{H}+0 \cdot 28 \mathrm{~W}+6 \cdot 88(\text { if male }) ; \\
R^{2} & =0 \cdot 93 \text { and } \mathrm{SEE}=2 \cdot 6 \mathrm{~kg}
\end{aligned}
$$

where $\mathrm{FFM}=$ fat-free mass (kg), $\mathrm{H}=$ height ( cm ) and $\mathrm{W}=$ weight $(\mathrm{kg})$.

In the development sample, the correlation between $\mathrm{FFM}_{\mathrm{ADP}}(\mathrm{kg})$ and height was $0.89(95 \%$ CI $0.85,0.92)$ and the correlation between $\mathrm{FFM}_{\mathrm{ADP}}(\mathrm{kg})$ and weight was 0.77 ( $95 \%$ CI $0.70,0.83$ ). Multicollinearity was noted (VIF $=2 \cdot 28$ ) and explained $93 \%$ of the total variance of FFM. Residuals had normal distribution and were homoscedastic. The model was parsimonious (Mallow's $C_{\mathrm{p}}$ coefficient $=4$ ).
Bland-Altman plots demonstrated good agreement without bias in the validation analyses. The limits of agreement were -6.0 to 6.2 kg (sD from mean of 3.1 kg ), indicating an acceptable validity. From the leave-one-out crossvalidation $\mathrm{SEE}=2.87$ and pseudo $R^{2}=0.91$ were obtained.

Using the validation sample, $\mathrm{FFM}_{\mathrm{EQ}}(\mathrm{kg})$ and $\% \mathrm{FM}_{\mathrm{EQ}}$ were calculated from the equation developed in the present study. A non-significant difference of $0 \cdot 12 \mathrm{~kg}$ resulted when $\mathrm{FFM}_{\mathrm{EQ}}(\mathrm{kg})$ was compared with $\mathrm{FFM}_{\mathrm{ADP}}(\mathrm{kg})$; in the same way, a non-significant difference of $0 \cdot 11 \%$ resulted when $\% \mathrm{FM}_{\mathrm{EQ}}$ was compared with $\% \mathrm{FM}_{\mathrm{ADP}}$ (Table 5). The mean of $\% \mathrm{FM}_{\mathrm{EQ}}$ for individuals with $\mathrm{BMI} \geq 30.0 \mathrm{~kg} / \mathrm{m}^{2}$ was $33.4 \%$ for men and $43.7 \%$ for women.

In the whole sample, prevalence of obesity was $74 \cdot 9 \%$ based on $\% \mathrm{FM}_{\text {ADP. }}$. The equation showed a sensitivity of $90 \cdot 3(95 \%$ CI $86 \cdot 8,93 \cdot 8) \%$ and a specificity of $58.0(95 \%$ CI $52 \cdot 1,63 \cdot 8) \%$ to identify obesity defined as excess of \% FM (Table 6).

Furthermore, the \%FM estimated had the ability to discriminate between people with obesity and people without obesity (Fig. 1). Ninety-three per cent and $65 \%$ of

Table 3 Classification of obesity according to fat mass percentage (\%FM) and BMI, by sex, in the sample of Mexican adults ( $n$ 275; 181 women) aged 20-63 years with BMI between 17.4 and $42.4 \mathrm{~kg} / \mathrm{m}^{2}$

|  | Women |  |  | Men |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | \% | $n$ |  | \% | $n$ |
| \%FM group |  |  | \%FM group |  |  |
| Underweight ( $\leq 20 \%$ ) | 1.7 | 3 | Underweight ( $\leq 10 \%$ ) | $3 \cdot 2$ | 3 |
| Normal (21-29 \%) | 6.1 | 11 | Normal (11-19\%) | 8.5 | 8 |
| Overweight (30-34 \%) | 14.9 | 27 | Overweight (20-24 \%) | $18 \cdot 1$ | 17 |
| Obese ( $\geq 35 \%$ ) | $77 \cdot 3$ | 140 | Obesity ( $\geq 25 \%$ ) | $70 \cdot 2$ | 66 |
| Total | $100 \cdot 0$ | 181 | Total | $100 \cdot 0$ | 94 |
| BMI group |  |  | BMI group |  |  |
| Underweight ( $<18.5 \mathrm{~kg} / \mathrm{m}^{2}$ ) | 1.7 | 3 | Underweight ( $<18.5 \mathrm{~kg} / \mathrm{m}^{2}$ ) | 0.0 | 0 |
| Normal (18.5-24.9 kg/m ${ }^{2}$ ) | 12.7 | 23 | Normal ( $18.5-24.9 \mathrm{~kg} / \mathrm{m}^{2}$ ) | 10.6 | 10 |
| Overweight ( $25.0-29.9 \mathrm{~kg} / \mathrm{m}^{2}$ ) | 45.9 | 83 | Overweight ( $25.0-29.9 \mathrm{~kg} / \mathrm{m}^{2}$ ) | 57.5 | 54 |
| Obese ( $\geq 30.0 \mathrm{~kg} / \mathrm{m}^{2}$ ) | 39.7 | 72 | Obesity ( $\geq 30.0 \mathrm{~kg} / \mathrm{m}^{2}$ ) | 31.9 | 30 |
| Total | $100 \cdot 0$ | 181 | Total | $100 \cdot 0$ | 94 |

Table 4 Models for estimation of fat-free mass using different predictive anthropometric variables in the sample of Mexican adults ( $n 275$; 181 women) aged 20-63 years with BMI between 17.4 and $42.4 \mathrm{~kg} / \mathrm{m}^{2}$

|  | Model 1 |  | Model 2 |  | Model 3 |  | Model 4 |  | Model 5 |  | Model 6 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Coef. | SE | Coef. | SE | Coef. | SE | Coef. | SE | Coef. | SE | Coef. | SE |
| Intercept | -46.03 | 5.67 | -42.51 | 6.04 | -40.7 | 6.04 | -34.67 | 6.88 | -25.33 | 8.19 | -19.33 | 8.92 |
| Weight (kg) | 0.28 | 0.02 | 0.30 | 0.02 | 0.31 | 0.02 | 0.35 | 0.04 | 0.42 | 0.05 | 0.41 | 0.05 |
| Height (cm) | 0.44 | 0.04 | 0.42 | 0.04 | 0.41 | 0.04 | 0.38 | 0.04 | $0 \cdot 33$ | 0.04 | 0.31 | 0.05 |
| Sex | 6.88 | 0.75 | 6.77 | 0.74 | 6.08 | 0.81 | 6.15 | 0.81 | 6.22 | 0.80 | 6.84 | 0.88 |
| SSF (mm) |  |  | -0.05* | 0.03 | -0.005* | 0.4 | -0.02* | 0.04 | -0.01* | 0.04 | -0.03* | 0.04 |
| TSF (mm) |  |  |  |  | -0.11 | 0.05 | -0.11 | 0.05 | -0.11 | 0.05 | -0.09* | 0.05 |
| SAD (cm) |  |  |  |  |  |  | -0.16* | 0.09 | -0.08* | 0.10 | -0.09* | 0.10 |
| WC (cm) |  |  |  |  |  |  |  |  | -0.10 | 0.05 | -0.08* | 0.04 |
| Age (years) | 0.93 |  |  |  |  |  |  |  |  |  | -0.04* | 0.05 |
| $R^{2}$ |  |  | 0.93 |  | 0.93 |  | 0.93 |  | 0.94 |  | 0.94 |  |
| SEE | $2 \cdot 643$ |  | $2 \cdot 627$ |  | 2.599 |  | 2.578 |  | 2.548 |  | 2.532 |  |
| VIF | $2 \cdot 28$ |  | $2 \cdot 17$ |  | 2.56 |  | 3.20 |  | 4.42 |  | 4.34 |  |
| MC | 4 |  | 5 |  | 6 |  | $132 \cdot 8$ |  | $600 \cdot 8$ |  | 1423.7 |  |

Coef., coefficient; SSF, subscapular skinfold; TSF, triceps skinfold; SAD, sagittal abdominal diameter; WC, waist circumference; Sex, 0 if a woman and 1 if a man; $R^{2}$, regression coefficient; SEE, standard sample error; VIF, variance inflation factor; MC, Mallow's $C_{p}$ coefficient. * $P<005$.

Table 5 Fat-free mass (FFM) and fat mass percentage (\%FM) estimated from the anthropometry-based equation and the air-displacement plethysmography (ADP) method in the validation sample ( $n$ 135) of Mexican adults aged 20-63 years with BMI between 17.4 and $42.4 \mathrm{~kg} / \mathrm{m}^{2}$

| Method | All ( $n$ 135) |  | Women ( $n$ 85) |  | Men ( $n 50$ ) |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Mean | SD | Mean | SD | Mean | SD |
| FFM (kg) |  |  |  |  |  |  |
| ADP (BOD POD ${ }^{\text {® }}$ ) | 48.47 | 9.7 | 42.56 | 5.3 | 58.52 | 6.6 |
| Equation | 48.59 | 9.2 | 42.76 | 4.7 | 58.50 | 5.8 |
| \%FM |  |  |  |  |  |  |
| ADP (BOD POD ${ }^{\circledR}$ ) | 34.41 | 9.2 | 38.40 | 7.1 | 27.61 | 8.4 |
| Equation | 33.30 | 8.0 | 38.22 | 5.9 | 27.66 | 6.5 |

people with and without obesity, respectively, were correctly classified.

## Discussion

While other studies have addressed the low specificity of the international \%FM cut-offs to identify metabolic
disorders ${ }^{(2)}$, to our knowledge, the present paper is the first to explore simple anthropometric variables to develop an equation to estimate $\% \mathrm{FM}$ in a sample of adults, in order to properly classify people with obesity.

Other studies did not consider simplicity of the predictor variables as a goal, possibly because they were not interested in using their results in population studies ${ }^{(18,19)}$ or

Table 6 Sensitivity, specificity, positive predictive value (PPV) and negative predictive value (NPV) analysis of BMI and the anthropometry-based equation in diagnosing obesity in the sample of Mexican adults ( $n 275 ; 181$ women) aged 20-63 years with BMI between 17.4 and $42.4 \mathrm{~kg} / \mathrm{m}^{2}$

| Validity test | BMI | $95 \% \mathrm{CI}$ | Equation | $95 \% \mathrm{CI}$ |
| :--- | :---: | :---: | :---: | :---: |
| Sensitivity (\%) | $54 \cdot 9$ | $49 \cdot 0,60 \cdot 7$ | $90 \cdot 3$ | $86 \cdot 8,93 \cdot 8$ |
| Specificity (\%) | $88 \cdot 4$ | $84 \cdot 6,92 \cdot 2$ | $58 \cdot 0$ | $52 \cdot 1,63 \cdot 8$ |
| PPV (\%) | $93 \cdot 4$ | $90 \cdot 5,96 \cdot 3$ | $86 \cdot 5$ | $82.5,90.6$ |
| NPV (\%) | $39 \cdot 6$ | $33 \cdot 8,45 \cdot 4$ | $66 \cdot 7$ | $61 \cdot 1,72 \cdot 2$ |

The cut-offs used to perform the validation of the \%FM equation were $25 \%$ for men and $35 \%$ for women; \%FM was obtained from airdisplacement plethysmography as gold standard.


Fig. 1 (colour online) Ability of fat mass percentage to discriminate between individuals with obesity ( $\triangle$ ) and without obesity (O), by sex (a, females; b, males), in the sample of Mexican adults ( $n$ 275; 181 women) aged 20-63 years with BMI between 17.4 and $42.4 \mathrm{~kg} / \mathrm{m}^{2}$; correspond to the $\% \mathrm{FM}$ cut-offs ( $\% \mathrm{FM} \mathrm{EQ}^{\mathrm{E}}$, fat mass percentage as estimated from the anthropometry-based equation; \%FM $\mathrm{MAD}_{\mathrm{ADP}}$, fat mass percentage as estimated from air-displacement plethysmography)
the adult population ${ }^{(23,33-35)}$; thus, they did not explore the ability of their equations to discriminate obesity.

Although anthropometry-based models for adult populations have been developed previously, usually they included different covariates, mainly skinfolds (Table 1). As we stated, one of the objectives of our study was to include the simplest variables with the lowest SEE and high predictive power. In 2007, Huerta et al. ${ }^{(23)}$ developed a model for the Mexican population of older adults, but that model used different covariates and different parameter values. One of the main contributions of our model is that, in addition to sex, it contains the same variables used to calculate BMI.

Based on BMI, the prevalence of obesity in our study is higher ( $37 \cdot 1 \%$ ) compared with that reported previously ${ }^{(6)}$ for the Mexican population ( $32 \cdot 4 \%$ ); the participation of volunteers probably explains this difference. On the other hand, prevalence of obesity based on $\% \mathrm{FM}$ compared with prevalence of obesity based on BMI is consistent with the results in other studies ${ }^{(12,36)}$.

We found a high correlation of $\% \mathrm{FM}_{\mathrm{ADP}}$ with BMI ( 0.82 for women and 0.77 for men), a result similar to that reported by other authors ${ }^{(4,37)}$. In order to choose the
simplest measurements, some variables highly correlated with others easier to measure were not selected in the final equation; for example, we found a high correlation of sagittal abdominal diameter with waist circumference $(0.82)$ and weight $(0 \cdot 80)$, results that are consistent with the literature ${ }^{(38)}$.

Some authors have found an association of age with body composition ${ }^{(39)}$; however, in the present study this variable had no significant effect on the FFM estimation although the range of ages included was wide. By contrast, this result is similar to that obtained by other authors interested in predicting $\mathrm{FFM}^{(40)}$.

As expected, height, weight and sex were highly correlated. Although we had presence of multicollinearity, VIF scores of less than 10 suggest that it was not a significant influence on the stability of the parameter estimated ${ }^{(41)}$.

In our study, the average $\% \mathrm{FM}_{\mathrm{ADP}}$ was higher for women than for men, similarly to previous studies ${ }^{(2,42)}$. Using data from 5100 Mexicans, Macias et al. ${ }^{(2)}$ found that the mean of $\%$ FM corresponding to $\mathrm{BMI} \geq 30.0 \mathrm{~kg} / \mathrm{m}^{2}$ was $49.3 \%$ for women and $35.8 \%$ for men; our results are similar for men ( $33.4 \%$ ) and lower for women
( $43.7 \%$ ). The difference could be explained by our sample size.

A cross-validation study of the new FFM prediction equation indicated high correlation with measured $\mathrm{FFM}_{\mathrm{ADP}}$; however, a slight tendency to underestimate $\mathrm{FFM}_{\mathrm{ADP}}$ was observed for both sexes. Similar tendency resulted from $\% \mathrm{FM}$. Other authors have found a tendency to overestimate $\% \mathrm{FM}$ at a lower $\% \mathrm{FM}$ and a tendency to underestimate $\% \mathrm{FM}$ at a higher $\% \mathrm{FM}^{(26)}$.

When we compare the root-mean-square error of our study it is smaller than that obtained in other studies. Furthermore, the variance explained by the FFM prediction equation obtained in our study $\left(R^{2}=0.93\right)$ was higher than that obtained by other authors, even when they used larger sample sizes. For example, the model developed for the Colombian population had two skinfolds, abdominal circumference, age and height as predictor variables ( $R^{2}=0.72$ ). The $R^{2}$ obtained in our study can be explained by the homogeneity of our sample, since we have more than $35 \%$ of the population with obesity. This suggests that this equation could fit populations with prevalence of obesity as high as our sample. In Mexico the prevalence of obesity is about $33 \%$.

The use of BMI as a measure of obesity can introduce misclassification problems that may result in important bias in estimating the effects related to obesity ${ }^{(43)}$. BMI had a high specificity, but a poor sensitivity to detect \%FM-defined obesity. Furthermore, the accuracy of BMI in diagnosing obesity is limited and it fails to discriminate between percentages of fat mass and lean mass in both sexes ${ }^{(44)}$. The sensitivity and specificity of BMI obtained in the present study were similar to those reported by other authors ${ }^{(12,44)}$.

However, one of the main reasons for continuing to use BMI and cut-off points proposed by the WHO is the comparability between populations; in that sense, the use of different indicators or cut-offs of obesity for each population could make this job harder. Another reason is the simplicity of its calculation.

Our findings also suggest that the magnitude of the obesity epidemic may be greater than that estimated by BMI. Using the gold standard definition of obesity as excess in \%FM, we show that the prevalence of obesity almost doubled from $39.7 \%$ using $\mathrm{BMI} \geq 30.0 \mathrm{~kg} / \mathrm{m}^{2}$ to $77.3 \%$ in women and from 31.9 to $70.2 \%$ in men. Differences in the obesity prevalence from both criteria are consistent with other studies ${ }^{(2,12,44)}$.

Some authors have suggested adjusting the cut-offs of BMI for obesity ${ }^{(42)}$; however, this proposal does not overcome the limitation of loss of comparability when the cut-off points depend on the population in question. Other authors have proposed to establish new international healthy body fat ranges ${ }^{(3)}$.

Potential limitations of our study include: (i) the limited generalization of our results to populations outside our sample, for example, to populations with lower prevalence of obesity; (ii) we used the two components
method as referent criterion although some authors say that it would be necessary to analyse at least three components to be a gold standard ${ }^{(45)}$, however, BOD POD was validated as a reference method ${ }^{(29,46)}$; and (iii) Siri's formula used in BOD POD was developed from a Caucasian population, 'Minnesota reference man'. This is a controversial point, because we used Siri's equation as a reference method to develop a model for a Mexican population.

Lee et al.'s ${ }^{(15)}$ finding suggests that the 'obesity paradox' controversy may be largely explained by low lean body mass, rather than low fat mass, in the lower range of BMI. In this sense, the FFM prediction equation developed in our study could help to estimate this component. This issue could not be addressed by the waist circumference measurement.

Finally, from our findings it is apparent that the diagnostic performance of the \%FM obtained from the new anthropometry-based prediction equation of FFM is a valid option mainly because of the ability of the \%FM estimated to discriminate between individuals with obesity and individuals without obesity.

## Conclusions

\%FM is an important physiological component and its estimation using a formula based on simple, direct and non-invasive anthropometric variables makes this kind of indirect method attractive to diagnose obesity.

Differences between ethnic groups are well established in the literature; however, this equation can be reliable to estimate fat mass in adults with similar physical characteristics to the participants of this study, and it can be used as a screening tool in clinical and epidemiological studies not only to estimate $\% \mathrm{FM}$ but also to discriminate the obese condition. Nevertheless, as we stated, with a non-representative sample it is difficult to generalize the results and further investigation into the findings described here needs to be undertaken.

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The lead author affirms that this manuscript is an honest, accurate and transparent account of the study being reported. The lead author affirms that no important aspects of the study have been omitted and that any discrepancies from the study as planned have been explained. Authorship: L.V.C.-P. and M.E.R.-R. participated in the planning, data acquisition, analysis, data interpretation and writing, of the study. J.V.-S. participated in the analysis, data interpretation and writing of the study. M.L.-C. participated directly in the planning, analysis, data interpretation and writing of the study. Ethics of human subject participation: This study was conducted according to the guidelines laid down in the Declaration of Helsinki and all procedures involving human subjects/patients were approved by Faculty of Medicine Ethics Committee from the National Autonomous University of Mexico. All participants provided written informed consent before participation.

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